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Endemicity of chytridiomycosis features pathogen over-dispersion

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32 Summary

33

34 1. Pathogens can be critical drivers of the abundance and distribution of wild
35 animal populations. The presence of an over-dispersed pathogen load
36 distribution between hosts (where few hosts harbor heavy parasite burdens
37 and light infections are common) can have an important stabilizing effect
38 on host-pathogen dynamics where infection intensity determines
39 pathogenicity. This may potentially lead to endemicity of an introduced
40 pathogen rather than extirpation of the host and/or pathogen.

41 2. Over-dispersed pathogen load distributions have rarely been considered in
42 wild animal populations as an important component of the infection
43 dynamics of microparasites such as bacteria, viruses, protozoa and fungi.

44 3. Here we examined the abundance, distribution and transmission of the
45 model fungal pathogen *Batrachochytrium dendrobatidis* (Bd, cause of
46 amphibian chytridiomycosis) between wild-caught *Litoria rheocola*
47 (common mist frogs) to investigate the effects of an over-dispersed
48 pathogen load distribution on the host population in the wild. We
49 quantified host survival, infection incidence and recovery probabilities
50 relative to infectious burden, and compared the results of models where
51 pathogen over-dispersion either was or was not considered an important
52 feature of host-pathogen dynamics.

53 4. We found the distribution of Bd load between hosts to be highly over-
54 dispersed. We found that host survival was related to infection burden, and
55 that accounting for pathogen over-dispersion allowed us to better
56 understand infection dynamics and their implications for disease control.
57 In addition, we found that the pattern of host infections and recoveries
58 varied markedly with season whereby (i) infections established more in
59 winter, consistent with temperature dependent effects on fungal growth,
60 and (ii) recoveries (loss of infection) occurred frequently in the field
61 throughout the year but were less likely in winter.

62 5. Our results suggest that pathogen over-dispersion is an important feature
63 of endemic chytridiomycosis, and that intensity of infection determines
64 disease impact. These findings have important implications for our

65 understanding of chytridiomycosis dynamics and the application of
66 management strategies for disease mitigation. We recommend quantifying
67 individual infectious burdens rather than infection state where possible in
68 microparasitic diseases.

69
70

71 **Key-words**

72

73 Aggregation, amphibian declines, frog, macroparasite, mark-recapture, microparasite,
74 multi-state, pathogen distribution, recovery, transition

75 **Introduction**

76

77 Pathogens are increasingly being identified as important drivers of the abundance and
78 distribution of wild animal populations (Altizer, Harvell & Friedle 2003; Grogan *et al.*
79 2014; Voyles *et al.* 2014). The complex host-pathogen dynamics that drive these
80 systems have classically been explained within the micro- and macro-parasite
81 epidemiological paradigm (the host is the relevant unit of study for microparasitic
82 infections while the individual parasite is the relevant unit of study for macroparasitic
83 infections). Within this paradigm, pathogens are categorised according to a number of
84 characteristics common to each group, such as the apparent degree of pathogen load
85 over-dispersion seen between hosts in macro- but not micro-parasites (Hudson *et al.*
86 2002; Wilson *et al.* 2002). Microparasites typically include viruses, bacteria, protozoa
87 and fungi, whereas macroparasites include larger organisms such as helminths and
88 arthropods. In this study we investigate and discuss the presence and importance of
89 pathogen over-dispersion as a key feature of endemic chytridiomycosis which has
90 classically been considered a microparasitic disease in terms of population effects.

91

92 Several recently emerged infectious pathogens (such as the fungus *Batrachochytrium*
93 *dendrobatidis* [Longcore, Pessier & Nichols 1999] in amphibians and Hendra virus in
94 bats; Wang *et al.* 1998) appear to defy clear categorisation within the aforementioned
95 paradigm as their dynamics fail to follow typical patterns (Briggs, Knapp &
96 Vredenburg 2010; Murray *et al.* 2013; Plowright *et al.* 2015). Furthermore, traditional
97 mitigation strategies based on the above paradigm are proving to be poor tools for the
98 control of these diseases in wild populations (for example, 'vaccination' by prior

99 exposure was unsuccessful in providing practical levels of protective immunity in the
100 study by Cashins *et al.* (2013)). An improved understanding of the dynamics of these
101 diseases is thus essential for managing them *in situ*, and will have broad applicability
102 to emerging infectious diseases in general. For example, recognizing the occurrence
103 of environmentally-associated seasonal peaks in chytridiomycosis severity (Phillott *et al.*
104 2013) led to the development of management strategies to address key
105 environmental factors and vulnerable life stages (using techniques such as head-
106 starting amphibians, translocations to habitats unfavourable to disease, and habitat
107 modification; for a review of strategies currently under development for mitigating Bd
108 *in situ*, see Scheele *et al.* (2014)).

109
110 *Batrachochytrium dendrobatidis* (hereafter Bd), the cause of the amphibian fungal
111 skin disease chytridiomycosis has had a devastating impact upon amphibian
112 populations around the world (through range contractions, population declines and
113 extirpations, and species extinctions; Skerratt *et al.* (2007)). In the past, single-celled
114 fungal pathogens like Bd have been typically considered as microparasites both
115 taxonomically and for the purposes of modeling their disease dynamics (Anderson &
116 May 1981). However, Bd infections demonstrate a number of features more common
117 to larger parasites such as helminths and arthropods (Hudson & Dobson 1998; Briggs,
118 Knapp & Vredenburg 2010).

119
120 Bd has a short life cycle within a single host involving two forms (infectious zoospore
121 and reproductive sporangium; Berger *et al.* 2005a). In contrast to typical
122 microparasites it appears to suppress an effective adaptive immune response in hosts
123 (Rosenblum *et al.* 2012; Cashins *et al.* 2013; Fites *et al.* 2013). While it is able to
124 multiply asexually at a moderate rate on individual hosts, duration of infection can be
125 long, and pathogenicity relies on high infectious burdens, a feature typical of
126 macroparasites (although severity may differ among individuals, populations and
127 species, with numerous intrinsic determinants of host susceptibility; Voyles *et al.*
128 2009; Vredenburg *et al.* 2010). Infectious burden also appears to be strongly
129 dependent on external factors affecting the life cycle of the pathogen, such as
130 temperature and moisture (Voyles *et al.* 2012), similar to macroparasitic diseases, and
131 hence population infections often display highly seasonal dynamics and

132 spatiotemporal distribution patterns consistent with environmental determinants
133 (Murray *et al.* 2013; Phillott *et al.* 2013).

134

135 Pathogen over-dispersion, another feature well recognised as common to
136 macroparasites, occurs with chytridiomycosis (Skerratt *et al.* 2011), but its underlying
137 causes and effects on disease dynamics have not been investigated. Pathogen over-
138 dispersion (otherwise known as parasite aggregation, pathogen aggregation, or
139 pathogen distribution heterogeneity) describes a distribution of infectious organisms
140 amongst hosts whereby most infected individuals have low infectious burdens, while
141 very few hosts have high burdens. Over-dispersion is an important characteristic for
142 understanding the dynamics of macroparasite diseases (Hudson & Dobson 1998), and
143 it implies that the infection intensity pattern (described by the intensity-frequency
144 curve) among hosts within a population tends to be highly positively skewed - thus
145 infectious organisms are spatially aggregated among hosts (Wilson *et al.* 2002).

146

147 Whilst it has been known that infectious burden is important in determining the
148 effects of chytridiomycosis on individuals (Voyles *et al.* 2009) the effects of the
149 distribution of infection burden, particularly over-dispersion, within populations has
150 largely been ignored (Briggs, Knapp & Vredenburg 2010; Vredenburg *et al.* 2010;
151 Skerratt *et al.* 2011). Over-dispersion likely provides one explanation for why some
152 species and populations persist with chytridiomycosis while others in previous studies
153 that do not exhibit over-dispersion have been driven to extinction (Lips *et al.* 2006;
154 Skerratt *et al.* 2007; Briggs, Knapp & Vredenburg 2010; Vredenburg *et al.* 2010). It
155 may also help to explain the inability to detect a difference in survival probability
156 between two disease states (infected and uninfected) in the multi-state mark-recapture
157 study by Briggs, Knapp and Vredenburg (2010), because the effects of a small
158 proportion of highly infected frogs may be unobserved when grouped with low
159 infection results.

160

161 Chytridiomycosis provides a unique opportunity to examine the phenomenon of
162 pathogen over-dispersion with a microparasite (Hudson & Dobson 1998; Skerratt *et*
163 *al.* 2011). Unlike typical microparasitic infections, the epidermal localization of
164 chytridiomycosis and the use of real time PCR enable the non-invasive and relative
165 quantification of burdens among hosts (Hyatt *et al.* 2007). Given pathogen over-

166 dispersion is likely a key feature of endemic chytridiomycosis, and infection intensity
167 affects both survival and infection transmission probabilities, examining its effects on
168 disease dynamics could have important implications for our understanding of the
169 disease ecology of microparasites.

170
171 We used the multi-state mark-recapture framework to investigate transmission and
172 recovery dynamics of endemic chytridiomycosis in *Litoria rheocola* (the common
173 mist frog; Liem 1974) in tropical north Queensland, Australia, as a function of
174 individual-level infection status, population-level apparent prevalence, and
175 environmental covariates. We chose this study species and tropical stream system as it
176 represents hundreds of similar species and systems that have survived introduction of
177 chytridiomycosis and now exist with endemic disease in Australia and the Americas
178 (Berger *et al.* 1998; Lips *et al.* 2006; Skerratt *et al.* 2007; Murray *et al.* 2009; Skerratt,
179 Speare & Berger 2011; Phillott *et al.* 2013). The study aimed to firstly, characterize
180 the presence of Bd pathogen over-dispersion in the context of a wild population of
181 endemically Bd infected amphibians, and secondly, to investigate infection and
182 recovery state transition dynamics throughout seasons and years. In particular, we
183 wanted to determine whether defining infection as a binary variable (two infection
184 states: uninfected and infected) or tertiary variable (three states: uninfected, and two
185 discrete levels of infectious load which takes into account pathogen over-dispersion)
186 affects our understanding of infection dynamics. Here we demonstrate that an over-
187 dispersed distribution of a microparasite within the host population plays an important
188 role in defining disease dynamics.

189

190

191 **Materials and methods**

192

193 **Species, site and sampling**

194 We collected mark-recapture encounter data (via toe-tip marks) for adult male
195 common mist frogs (*Litoria rheocola*) from a 150 m stream transect in lowland
196 tropical rainforest of Tully Gorge National Park (145° 38' E 17°46' S, 130 m above
197 sea level), Queensland, Australia over 22 trips between November 2005 and October
198 2007 (see Phillott *et al.* 2013 for further details of field work at this site). Bd is

199 suspected to have arrived at this site around 1989. Although annual survival rates are
200 low (12%) there is high recruitment (91%) and the population appears stable (Phillott
201 *et al.* 2013). *L. rheocola* is an obligate stream-breeder, and the breeding season for
202 this population occurs from May to August (coinciding with the dry winter season;
203 Bureau of Meteorology 2008) however adult males maintain calling territories at the
204 stream throughout the year (Hodgkison & Hero 2002; Phillott *et al.* 2013). Individual
205 frogs were skin-swabbed at every capture (maintaining strict hygiene, and following
206 standard protocols; Phillott *et al.* 2010; Phillott *et al.* 2013), and swabs were analyzed
207 for the presence of Bd DNA via quantitative PCR (qPCR; one well, one zoospore
208 equivalent [*zse*] considered positive; Hyatt *et al.* 2007; Skerratt *et al.* 2011).

209

210 **Multi-state mark-recapture modeling**

211 We applied the information theoretic approach using Akaike's Information Criterion
212 (IT-AIC, following the steps outlined in Phillott *et al.* 2013) to explore state-specific
213 endemic chytridiomycosis infection dynamics using the multi-state mark-recapture
214 (MSMR) framework (see Appendix S1 for more details). We hence performed two-
215 and three-state multi-state modeling with program MARK (version 6.0; White,
216 Kendall & Barker 2006) to elucidate the individual-level effect of chytridiomycosis
217 infection on survival probability in the field by assigning frogs to an infection state at
218 each capture via qPCR results. We particularly wanted to determine the probabilities
219 for infection and recovery transitions, in order to understand the nature of infection
220 dynamics *in situ* (Murray *et al.* 2009; Cooch *et al.* 2012).

221

222 We investigated the best predictors (several, due to model uncertainty) from the
223 Cormack-Jolly-Seber (CJS) analysis for existing survival and recapture parameters for
224 this dataset (Phillott *et al.* 2013) in the new context of state transition probabilities.
225 Hence we investigated survival as a function of infection status (γ), apparent trip
226 prevalence (π), mean daily maximum temperature ($^{\circ}\text{C}$) for the 28 days preceding each
227 trip (λ), and a cyclical seasonal linear trend variable (τ , where autumn is considered
228 equivalent to spring). Recapture probability was investigated as a function of infection
229 status (γ), mean daily relative humidity (%) at maximum temperature for the 28 days
230 preceding each trip (ϵ), mean daily radiation (MJ/m^2) over 28 days preceding each trip
231 (ζ), and capture effort (in days per trip δ). Weather variables were obtained from the

232 SILO climate database which provides spatially interpolated values from regional
233 meteorological stations (Jeffrey *et al.* 2001; Bureau of Meteorology 2008).

234

235 We defined infection status (γ) as a time-varying individual covariate categorized into
236 either two or three states on the basis of infection intensity (z_{se}) at each capture. In the
237 two-state analysis, A = Bd negative (uninfected) and B = Bd positive (infected). In the
238 three-state analysis, Bd load was discretized into groups: A = Bd negative, B = 1-4 z_{se}
239 “low”, C > 4 z_{se} “high”. This low-burden group of hosts is the most poorly defined in
240 terms of disease processes; individuals may be newly infected, recovering, resistant,
241 their burdens may represent background contamination, or they may contain
242 unaccounted sampling or laboratory error (McClintock *et al.* 2010). The chosen
243 threshold between infection states (4 z_{se}) allowed us to separately model the
244 transmission dynamics of this low-burden group and eliminated potential confounding
245 from the high-infected host group. In addition, multi-state analysis methods have high
246 data requirements, and this threshold permitted Bd positive results to be split evenly
247 between states B (low intensity) and C (higher intensity) providing sufficient power
248 for analysis (66 samples $z_{se} \leq 4$; 64 samples $z_{se} > 4$; Fig. 1). We acknowledge that by
249 artificially discretizing the continuous variable z_{se} into low and high categories of
250 intensity of infection there is some loss of information and some potential
251 misclassification of infection levels close to the cutoff value (although the diagnostic
252 accuracy of the quantitative PCR at James Cook University is very high, sensitivity
253 73% and specificity approximating 100%; Hyatt *et al.* 2007; Skerratt *et al.* 2011).
254 Regardless, the results remain interpretable in terms of the effects of comparative
255 levels of infection. The sample size was not sufficient for categorization into
256 additional levels of infection intensity such as a moderate group, and although it
257 would have been ideal to perform sensitivity analyses for determining this threshold,
258 unfortunately these were not possible due to data sparseness. Indeed, varying the cut-
259 off value by a single zoospore in either direction rendered the data unanalysable due
260 to failure of numerical convergence of the majority of models tested.

261

262 The state transition parameter ψ_i^{rs} defines the probability that an individual in state r
263 at time i will be in state s at time $i + 1$. Importantly where there are more than two
264 states, this includes the probability of transitions from each state in the MSMR Jolly-
265 Movement Model (JMV; Lebreton *et al.* 2009), including the probability of remaining

266 in the same state ψ_i^{rr} , and the outgoing probabilities for each state must sum to one
267 (Fig. 1). States in this study represent discrete infection conditions (defined by zse
268 infection intensities) in which the marked individual may potentially be encountered,
269 conditional on being in that state and alive. Following the results in Phillott *et al.*
270 (2013), and to incorporate both individual and population-level effects, we
271 hypothesized that state transition probabilities are influenced by infection status (γ),
272 apparent trip prevalence (π) and seasonal environmental covariates such as
273 temperature (λ). As an example of how these effects might influence the transitions
274 between states, recoveries should be associated with increased ambient temperature to
275 reduce Bd growth (Voyles *et al.* 2012) and promote host thermoregulatory
276 immunomodulation (Richards-Zawacki 2010). Similarly, recoveries should also be
277 associated with reduced prevalence as they require an absence of re-infection.
278
279 We applied goodness of fit tests and applied the most conservative estimate of $\hat{c} =$
280 1.110 and 1.097 for the two and three-state analyses respectively (Appendix S1).
281 Candidate model sets for two and three-state analyses were constructed separately *a*
282 *priori* using a restricted form of the all subsets approach, and tested systematically
283 (Appendix S1; Lukacs, Burnham & Anderson 2010; Hegyi & Garamszegi 2011;
284 Doherty, White & Burnham 2012). We constructed models using the intercept design
285 matrix coding format and the logistic (logit) link function. Where numerical
286 convergence was suspect, we employed the alternate optimization routine from within
287 MARK, and assessed each model individually for estimable parameter count,
288 adjusting as necessary (Lebreton *et al.* 2009; Cooch *et al.* 2012). We used QAIC_c to
289 rank model parsimony (Burnham & Anderson 2002), model averaging to reduce
290 selection bias (Lukacs, Burnham & Anderson 2010), and we estimated monthly
291 parameter probabilities ($1 \text{ month} = \frac{365}{12} \approx 30.42 \text{ days}$), reporting unconditional 95%
292 confidence intervals (95% CI; Burnham & Anderson 2002). Akaike weights were
293 used to determine relative variable importance from entire candidate model sets
294 (Doherty, White & Burnham 2012), and we report evidence ratios and model
295 averaged effect sizes where appropriate for comparisons between states (Burnham &
296 Anderson 2002). Model averaged effect sizes were based on model averaged real
297 parameter estimates and confidence intervals were unbounded on the real probability
298 scale using the delta method for difference between two variances with the model

299 averaged variance-covariance matrix. Raw encounter history and predictor variable
300 data together with model averaged parameter estimates for the three-state analysis are
301 available in Appendix S3. We additionally performed a discrete time simulation for a
302 population of adult frogs employing the model-averaged trip-based parameter
303 estimates from the three-state multi-state mark-recapture analysis over the study
304 period to demonstrate the impact of estimated state transition and survival parameters
305 on actual population numbers. Detailed methods and results from this simulation are
306 available in Appendix S1 and S2.

307

308

309 **Results**

310

311 **Infection pattern summary**

312 We made 424 captures of 243 uniquely marked adult male *L. rheocola* frogs
313 throughout the two year study period (109 frogs were caught more than once). Forty-
314 seven frogs (43% of those caught more than once) changed infection state at least
315 once (became infected or recovered), and 13 frogs (28% of those caught more than
316 twice) changed state two or more times (although only three of these, 23%, were re-
317 infected after recovery). State transitions were approximately even with 28 infection
318 and 34 recovery transitions. Two frogs gained and lost infection several times. The
319 highest infection intensity recorded prior to recovery was 123 *zse*. Apparent Bd
320 infection prevalence for the whole study period was $130/421 = 0.3088$ (binomial 95%
321 CI 0.2650 to 0.3553 by Clopper-Pearson method, assuming statistical independence).
322 The intensity-frequency histogram for qPCR swab results for the whole study period
323 was highly positively skewed (Fig. 2; 291 records for Bd negative and 21 high *zse*
324 records were truncated for visualization; $N = 421$, range 0 to 4028 *zse*). The variance
325 to mean ratio of infectious organisms per host (s^2/m) was 2227.47 (very much higher
326 than one, indicative of pathogen over-dispersion). To further examine evidence of
327 over-dispersion, the Weibull distribution ($\alpha = 0.46901$, $\beta = 15.259$) and negative
328 binomial distribution were fit to the population-level apparent infection prevalence
329 data (Fig. 2), and the corrected moment estimate of k (of the negative binomial
330 distribution) was 0.0069, indicating a high degree of pathogen over-dispersion
331 (Wilson *et al.* 2002). Testing the most parsimonious zero-inflated negative binomial

332 model (with season and year as covariates) against the standard negative binomial
333 model indicated that there was no significant difference between the models (AIC-
334 corrected Vuong z-statistic = 0.531, $p = 0.298$), and thus zero inflation (excessive
335 ‘uninfected’ samples) is unlikely to have contributed to the observed over-dispersion
336 (Zuur *et al.* 2009; see Appendix S2 for further details).

337

338 **Multi-state mark-recapture results**

339 Model averaged parameter estimates using individual time-varying infection intensity
340 results revealed marked seasonality in survival and transition probabilities in both
341 analyses (for example in the three state analysis, survival probability varied seasonally
342 from 0.48 in winter to 0.79 in summer for frogs in the high burden state). Monthly
343 model averaged estimates for state-dependent survival, recapture and state transition
344 probabilities are reported with unconditional 95% confidence intervals in Figs 3 and 4
345 for two- and three-state multi-state analyses, respectively (see Appendix S2 for ranked
346 tables of model results). Although the two and three-state MSMR analyses cannot be
347 directly compared with AIC due to their differing candidate model sets, the three state
348 MSMR analysis better captured the over-dispersion of infectious burden because it
349 treated relatively highly infected individuals independently to low and uninfected
350 individuals. In contrast, the two state analysis grouped all infected individuals
351 together regardless of infectious burden. While survival differed between infected and
352 uninfected frogs in the two-state analysis, apparent survival probability estimates for
353 the infected group were higher than those for the uninfected group, except during one
354 winter trip session. Confidence intervals for the infected group were considerably
355 wider, however, and overlapped those for the uninfected group for all trip sessions
356 (Fig. 3a). In comparison, when the infected group was separated into two infection
357 categories (group B with 1-4 z_{se} , group C $> 4 z_{se}$) in the three-state analysis, frogs
358 with differing levels of infectious burden had differing survival probabilities (frogs
359 with $> 4 z_{se}$ had consistently lower survival; Fig. 4a). While recapture probabilities
360 were relatively stable throughout the study period in both analyses, in the two-state
361 analysis both uninfected and infected frogs had similar recapture probabilities (Fig.
362 3b), whereas in the three-state analysis the low-burden group had low recaptures
363 compared with the high-burden group (although confidence margins were wide in the
364 three-state analysis; Fig. 4b, see Appendix S2 for further details).

365

366 Parameter estimates revealed marked seasonality in state transition probabilities
367 between infection states. In the two-state analysis, frogs were much more likely to
368 become infected in winter (correlating with prevalence), while there was a moderate
369 reduction in the probability for recovery transitions during this period in the infected
370 group (Fig. 3c). The three-state analysis further highlighted these trends with some
371 exceptions despite overlapping confidence margins (transitions constituting the gain
372 of or increase in infectious burden shown in Fig. 4c; reduction of infectious load or
373 loss of infection transitions shown in Fig. 4d). The highest probability for infection
374 transitions occurred during winter from the uninfected (group A) to low-burden frogs
375 (group B). Recovery transition (loss of infection) probabilities were seasonal, peaking
376 during summer and autumn, and were similar between both high and low-burden
377 groups. Stationary transition probabilities (shown in Fig. S1, Appendix S2) were
378 derived from the aforementioned model-averaged transition probabilities and
379 probability theory which states that the sum of the probabilities of leaving each state
380 must equal one. Hence throughout most of the year, among those surviving a
381 sampling interval, frogs were most likely to either remain in the uninfected state, or
382 return to that state through infection recovery (Fig. S1, Appendix S2). Low-burden
383 frogs (group B) were observed to increase their infectious load (to group C) at a
384 relatively low and stable rate throughout the study (Figs 4c, S1b, Appendix S2).
385 Hypothetical population dynamics (including variation in total population size) based
386 on these transition and survival probabilities are exemplified in a series of three
387 population dynamics simulation models illustrated in Fig. 5.
388
389 Despite model selection uncertainty, the most parsimonious models in both analyses
390 modeled apparent survival and state transition probabilities as a function of a
391 multiplicative interaction between individual-level infection state and population-level
392 infection prevalence (the models $S(\gamma \times \pi)\rho(\delta)\psi(\gamma \times \pi)$ and $S(\gamma \times \pi)\rho(\gamma +$
393 $\zeta)\psi(6\gamma \times \pi)$, with 9.1% and 20.4% support for two- and three-state analyses,
394 respectively). Ranked relative predictor variable importance (reporting only those
395 >0.1) for the two-state analysis were prevalence ($\pi = 0.6228$) and temperature ($\lambda =$
396 0.3400) for survival; days ($\delta = 0.36301$), radiation ($\zeta = 0.29044$) and relative
397 humidity ($\epsilon = 0.24491$) for recapture; and prevalence ($\pi = 0.9254$) for transition. For
398 the three-state analysis these were prevalence ($\pi = 0.8700$) and temperature ($\lambda =$

399 0.1298) for survival; radiation ($\zeta = 0.5109$), relative humidity ($\varepsilon = 0.2632$) and days (δ
400 $= 0.1674$) for recapture; and prevalence ($\pi = 0.9264$) for transition.

401

402 The model averaged effect size as a mean across trips for the survival difference
403 between infected and uninfected groups in the two-state analysis was 0.1070, with the
404 infected group demonstrating higher apparent survival overall (95% CI -0.0577 to
405 0.2717). Similarly, the model averaged effect sizes for survival in the three-state
406 analysis were as follows: B-A 0.1184 (95% CI -0.0448 to 0.2815), B-C 0.2610 (95%
407 CI -0.1573 to 0.6794) and A-C 0.1427 (95% CI -0.2086 to 0.4940). While there was
408 limited support for an effect of individual infection status on apparent survival in the
409 two-state analysis (the evidence ratio comparing most parsimonious models with and
410 without γ was 2.8222), there was correspondingly strong support in the three-state
411 analysis (evidence ratio 695.20), and strong support in both analyses for an effect of
412 infection status on state transition probability (evidence ratios > 918.90 and 319.36 for
413 the two- and three-state analyses, respectively; Lukacs *et al.* 2007).

414 **Discussion**

415

416 **Pathogen over-dispersion and survival probabilities**

417 We found marked aggregation of Bd within our endemically infected wild amphibian
418 population, as demonstrated by a highly over-dispersed intensity-frequency
419 distribution curve (Fig. 2). Thus, while most infected individuals had low burdens, a
420 few hosts had high burdens. Categorizing infectious burdens into low or high groups
421 based on qPCR swab results allowed us to partially resolve paradoxical results from
422 our two-state analysis which were similar to those reported by Briggs, Knapp and
423 Vredenburg (2010). The model averaged estimates from our two-state analysis
424 incongruously revealed a lower apparent survival probability for uninfected frogs
425 compared with infected frogs, although confidence intervals for the infected state
426 were wide (Fig. 3a). After taking pathogen over-dispersion into account, apparent
427 survival probability of infected frogs fell to either side of the uninfected group, with
428 high-burden frogs having the lowest survival estimates (Fig. 4a). The reason for a
429 difference between the two and three state analyses is the high degree of pathogen
430 over-dispersion and its differential effects; approximately half the infected frogs were
431 classed in the low-burden group (Fig. 2). In addition, infection intensity was found to

432 be seasonally associated with survival as well as transmission and recovery
433 probabilities. Our results are consistent with previous field work showing over-
434 dispersion (Skerratt *et al.* 2011), and linking reduced survival with higher Bd
435 infection intensities (Murray *et al.* 2009), and also demonstrates that quantifying
436 infectious burdens is key to understanding the ecology of chytridiomycosis.

437

438 We used the MSMR framework to provide dynamic estimates of first-order Markov
439 infection state transition probabilities and state-dependent survival estimates from
440 field data whilst accounting for imperfect detection (Murray *et al.* 2009; Cooch *et al.*
441 2012). Compared with the single-state Cormack-Jolly-Seber model (Phillott *et al.*
442 2013), the MSMR framework permits reassessment of individual disease status at
443 each capture, which is essential for examining individual-level infection dynamics and
444 survival probabilities in a system where infection status fluctuates. Most disease
445 studies utilizing MSMR analyses to date have categorized individuals on the basis of
446 their infection status (uninfected versus infected states; Murray *et al.* 2009; Briggs,
447 Knapp & Vredenburg 2010). This binary definition in the presence of pathogen over-
448 dispersion can greatly diminish our understanding of survival and state transition
449 probabilities, and here we demonstrate the importance of this effect through
450 comparisons of two and three state analyses.

451

452 **Transition probabilities – infection and recovery**

453 In our study, frogs gained and lost infection frequently, consistent with previous field
454 data on mountain yellow-legged frogs (*Rana muscosa* and *R. sierrae*) in temperate
455 USA (Briggs, Knapp & Vredenburg 2010), and some individuals demonstrated
456 numerous state transitions. Comparing two- and three-state analyses helped resolve
457 the nature and magnitude of transition probabilities between disease states (Figs 3c, 4c
458 and 4d). As expected from previous studies on the temperature dependence of Bd
459 (Voyles *et al.* 2012), we found that frogs were most likely to become infected during
460 winter months (June to August in the southern hemisphere), with the transition to a
461 low infectious burden (1-4 *zse*) being most probable (Fig. S1a, Appendix S2).

462 Alternatively, recovery from both low and high infectious burdens was equally
463 probable and high throughout most of the year, dropping moderately during winter
464 (Figs S1b and S1c, Appendix S2).

465

466 A relatively long incubation period (roughly 3-8 weeks between exposure and clinical
467 signs; Berger *et al.* 2005b; Voyles *et al.* 2009) in an environmentally responsive
468 pathogen means more chance for pathogen-adverse environmental conditions (such as
469 temperature spikes) to favour host recovery and survival. Thus, recovery transitions
470 may be favoured over infection transitions throughout most of the year in areas with
471 higher temperatures such as at low elevation tropical regions. A long incubation
472 period also artificially increases point prevalence measures and decreases mortality
473 measures in comparison with pathogens that have short incubation periods. This is
474 also likely to lead to a highly over-dispersed intensity-frequency distribution because
475 most of the infected population is in the subclinical phase of the disease at any point
476 in time (in endemically infected populations, unlike propagating epidemics which can
477 rapidly lead to widespread mortality). Re-infection transitions were comparatively
478 uncommon, however (only three of the 13 frogs that were observed to change state
479 two or more times), possibly suggesting adaptive immunity may occur in the field.
480 However, the third simulation scenario (Fig. 5c) assumed no effect of adaptive
481 immunity and resulted in population dynamics consistent with our expectations and
482 dynamics observed in the wild.

483

484 **Uninfected frogs versus those with low infectious burdens**

485 Finding that uninfected frogs had lower apparent survival probabilities than those in
486 the low-burden state (Fig. 4a) was unexpected. The difference in apparent survival
487 between these infection states was small to moderate (11% for two-state analysis, 12-
488 26% for three-state analysis). Perhaps this survival discrepancy, and part of the cause
489 for the high level of pathogen over-dispersion, is due to the low infection group
490 representing a greater proportion of more resistant individuals. Under laboratory
491 conditions conducive to disease progression, infections occur only at low levels for
492 about a week post-exposure (Hyatt *et al.* 2007) suggesting low infections in
493 susceptible wild individuals would only be maintained if conditions for the disease
494 were suboptimal or if individual frogs were relatively resistant. In comparison, the
495 uninfected group would contain susceptible individuals that eventually become
496 exposed and die from the disease but are not re-caught prior to death. Similarly, the
497 lower winter survival probabilities in the low burden and uninfected frogs compared
498 with other seasons (given that pathogenicity relies on high infectious burdens, and that
499 these burdens are driven by weather; Voyles *et al.* 2009; Murray *et al.* 2013) is likely

500 due to frogs increasing their intensity of infection and dying without being detected in
501 the high burden state.

502

503 Alternatively, the above discrepancy may be due to emigration confounding in
504 capture-mark-recapture studies (Murray *et al.* 2010; Schmidt 2010). For example,
505 differential permanent emigration rates between the two states may lead to different
506 apparent survival probabilities. We have no *a priori* reason to suspect higher
507 emigration in uninfected frogs or in those with high zoospore burdens compared with
508 those having low burdens (Roznik *et al.* 2015). Rather, frogs appeared to maintain
509 calling territories on the stream year-round suggesting site fidelity (Phillott *et al.*
510 2013). Tracking studies comparing movements of uninfected and infected frogs could
511 be used to resolve potential confounding due to emigration if it appears to be an issue
512 (Roznik & Alford 2015).

513 **Implications of pathogen over-dispersion**

514 Implications of pathogen over-dispersion for the management of endemic
515 chytridiomycosis can be separated into two categories; those that affect the way we
516 study, model and report this disease; and those that affect actual disease dynamics. In
517 the first instance, we have demonstrated empirically that failing to account for
518 different levels of infectious burden between hosts can lead to biases in understanding
519 of population dynamics (for example, through mark-recapture state categorizations, or
520 ecological modeling). In addition, we highlight that the commonly reported measure
521 of disease abundance, population infection prevalence, should not be used to compare
522 populations with differing levels of pathogen over-dispersion. In the second instance,
523 pathogen over-dispersion impacts population dynamics where infectious burden
524 affects 1) pathogenicity, 2) the rate of production of the infectious stage released to
525 the environment, or 3) the degree of host resistance or immunity (May & Anderson
526 1979). The first two conditions occur in chytridiomycosis, based on this study and
527 past work (Hyatt *et al.* 2007; Murray *et al.* 2009).

528

529 The specific effects of pathogen over-dispersion on populations will likely depend on
530 the degree and predominant causes of the observed over-dispersion, and elucidating
531 these may assist with predicting long-term population outcomes and hence
532 management approaches. There are three main potential causes of observed over-
533 dispersion including 1) heterogeneous exposure, 2) variable multiplication within the

534 host, and 3) sampling artifact (Hudson & Dobson 1998). In the context of endemic
535 chytridiomycosis the first two are likely to be most important. To rule out the third
536 option, we fitted and compared mixture models and a standard negative binomial
537 model to the infection intensity data. We demonstrated that over-dispersion is an
538 important component of the variance structure of the system and that it could not be
539 explained by zero-inflation, which largely rules out sampling artifact as a major
540 contributor. Our results demonstrate that environmental variation (seasonality) was a
541 significant predictor of the underlying infection process (observed in both the mark-
542 recapture analyses and the count model component in the zero-inflated negative
543 binomial model, Appendix S2). This result is consistent with previous studies
544 (Murray *et al.* 2013). Variable multiplication in the host due to environmental
545 seasonality is likely to be an important cause of over-dispersion in endemic
546 chytridiomycosis (Berger *et al.* 2009). This is also likely to lead to seasonal
547 heterogeneity in exposure as zoospores can either disperse or re-infect the same host.

548

549 In our study, prevalence was ranked with highest predictor variable importance in the
550 three state analysis for both survival and transition probabilities, despite a strong
551 correlation with temperature, suggesting that a host component also plays an
552 important role in determining the distribution of infectious burdens. We were
553 unfortunately unable to examine host immunity/resistance as a direct covariate in
554 either the mark-recapture analyses or the mixture models for over-dispersion as there
555 is currently insufficient knowledge of resistance markers to chytridiomycosis in
556 amphibians (although this is an active area of research).

557

558 Two major reasons Bd infection dynamics differ from the classic microparasite
559 paradigm are 1) the relative absence of effective adaptive immunity to Bd (Cashins *et al.*
560 *et al.* 2013; Fites *et al.* 2013; McMahon *et al.* 2014) and the prominence of variation in
561 host susceptibility to infection (Berger *et al.* 2005b; Tobler & Schmidt 2010), and 2)
562 the ectothermic nature of amphibian hosts leading to variations in pathogen growth
563 due to thermal suitability (Murray *et al.* 2013). The former may be associated
564 suppression and evasion of the host immune system, as well as host physiological and
565 behavioural characteristics (Tobler & Schmidt 2010; Savage & Zamudio 2011).
566 Species that have been driven to extinction by chytridiomycosis appear to be highly
567 susceptible, have little variation in host susceptibility (Carey *et al.* 2006; Skerratt *et*

568 *al.* 2007; Berger *et al.* 2009; Bataille *et al.* 2013), and have occurred in areas that were
569 highly favourable for the pathogen (Murray *et al.* 2011). Thus, it is important to
570 identify the predominant cause of pathogen over-dispersion as this may provide an
571 indication of potential long-term persistence of the population.

572

573 **Conclusions**

574 In conclusion, we have shown that pathogen over-dispersion is an important feature of
575 a microparasitic disease, in this case endemic chytridiomycosis. Overlooking non-
576 uniform pathogen distributions in microparasitic diseases may lead to paradoxical
577 interpretations of disease dynamics. We also show that Bd infections occur seasonally
578 and that recoveries are common and likely important for population persistence.
579 Future management of endemic chytridiomycosis might focus on environmental
580 manipulation to favor host recoveries (Scheele *et al.* 2014). Understanding the main
581 causes of pathogen over-dispersion will indicate whether other disease control
582 interventions should be targeted predominantly towards 1) assisting the longer-term
583 evolution of resistance within the population via selection techniques, 2) reducing
584 exposure and transmission of infection between hosts, 3) bolstering population size
585 through approaches directed at habitat conservation, or 4) minimizing other
586 threatening processes. We recommend quantifying individual infectious burdens
587 rather than infection state where possible in microparasitic diseases.

588

589 **Data Accessibility**

590

591 Raw encounter history and predictor variable data, together with three-state model
592 averaged parameter estimates are available in Appendix S3 online.

593

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595

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605

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812

813 **Figure legends**

814 **Figure 1.** Example schematic illustrating state transition probabilities (ψ) and survival
815 probabilities (S) for the respective infection states at capture session six (drawn from
816 the three-state multi-state analysis). The notation ψ^{rs} indicates the monthly state
817 transition probability from state r to state s from time (capture session) i to $i+1$, and S^t
818 represents survival probability from time i to time $i+1$, for individuals in state t . Circle
819 sizes are representative of the relative expected population size (from the simulation),
820 and arrow line thicknesses represent the relative magnitude of the respective
821 probabilities.

822

823 **Figure 2.** Intensity-frequency histogram showing highly over-dispersed distribution
824 of infectious organisms between individual hosts (highly positively skewed), together
825 with the fitted Weibull and negative binomial distributions. $N = 421$; 291 Bd negative
826 records and 21 high zoospore records were truncated for visualization; original data
827 range 0 to 4028.

828

829 **Figure 3.** Model averaged estimates for monthly (a) survival probability, (b) recapture
830 probability, and (c) state transition probability with unconditional 95% confidence
831 intervals from the two-state multi-state analysis for male adult *L. rheocola* at Tully.

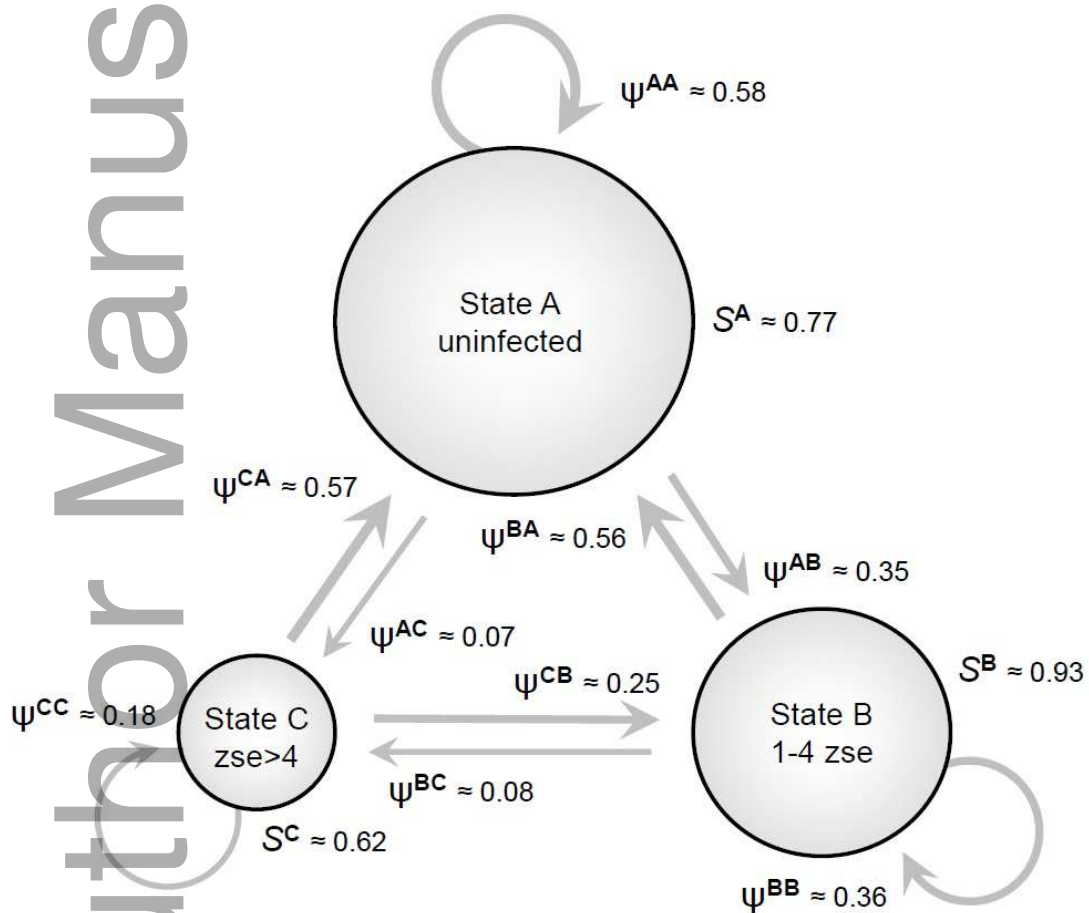
832

833 **Figure 4.** Model averaged estimates for monthly (a) survival probability, (b) recapture
 834 probability, (c) infection transition probabilities, and (d) recovery transition
 835 probabilities with unconditional 95% confidence intervals from the three-state multi-
 836 state analysis for male adult *L. rheocola* at Tully. States are defined as: state A = Bd
 837 negative (uninfected), state B = 1-4 zse, state C >4 zse.

838 **Figure 5.** Outcomes from the demographic simulation without recruitment (a); with
 839 recruitment to force population stability (b); with recruitment as model-averaged
 840 estimates from Phillott *et al.* (2013) (c).

841

842 **Figure 1.**

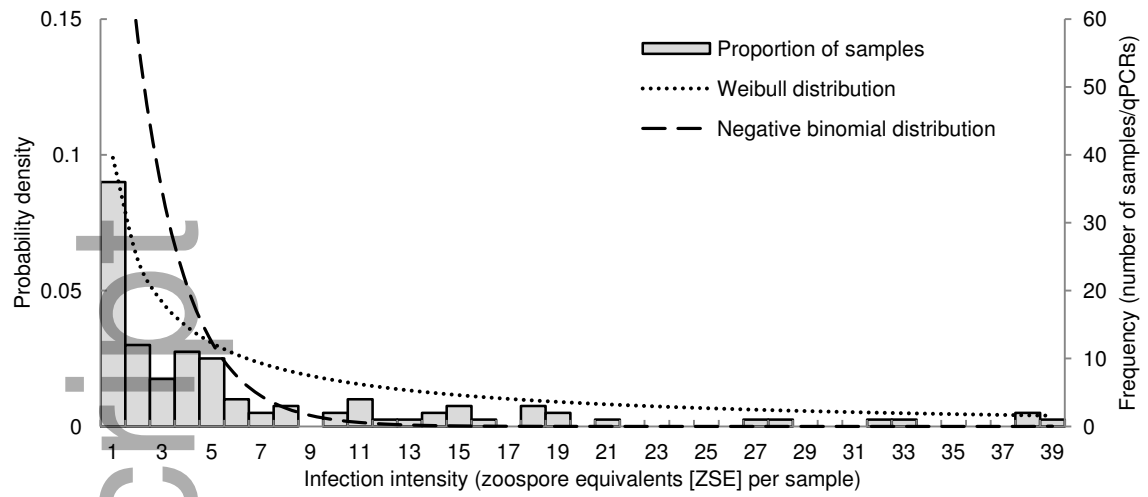


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844

845 **Figure 2.**

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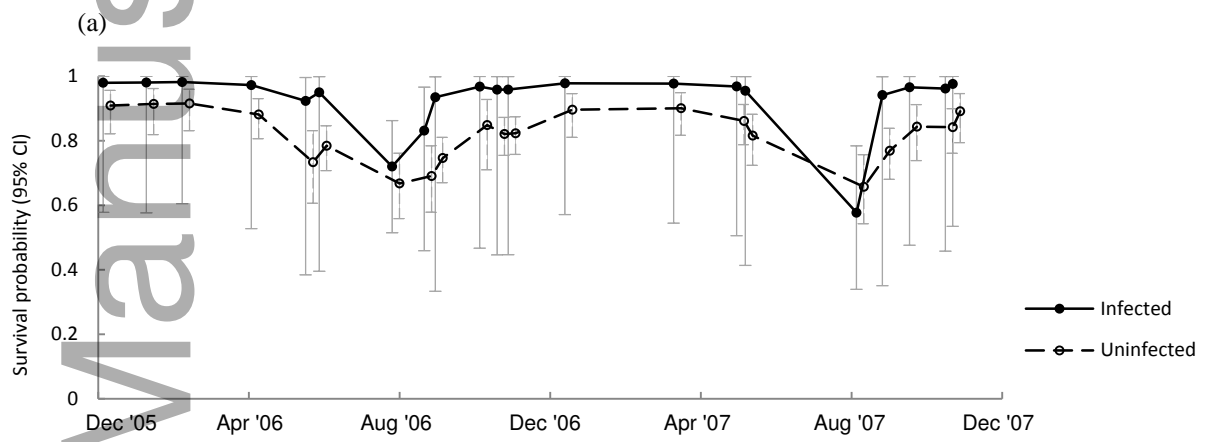


847

848 **Figure 3.**

849

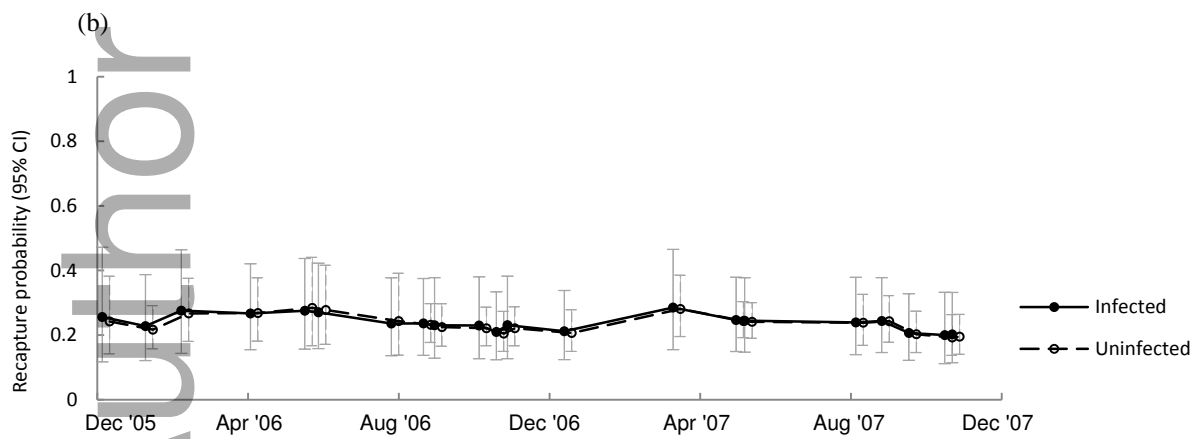
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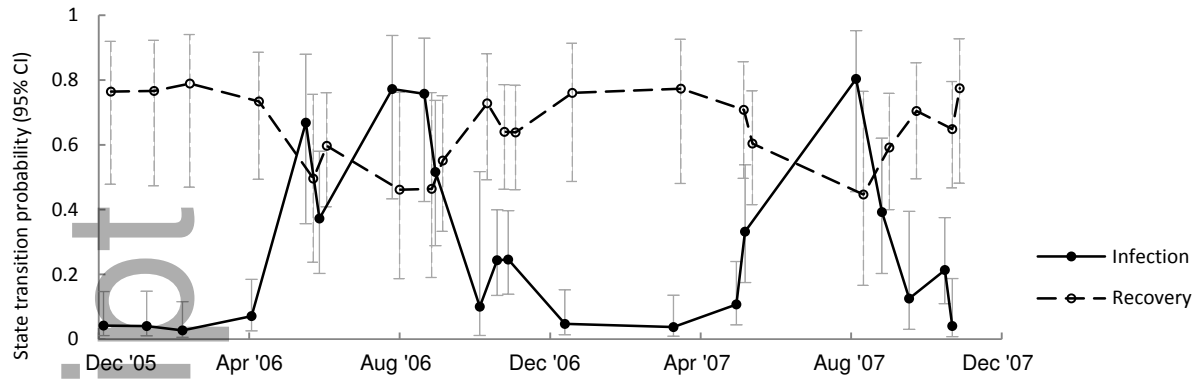


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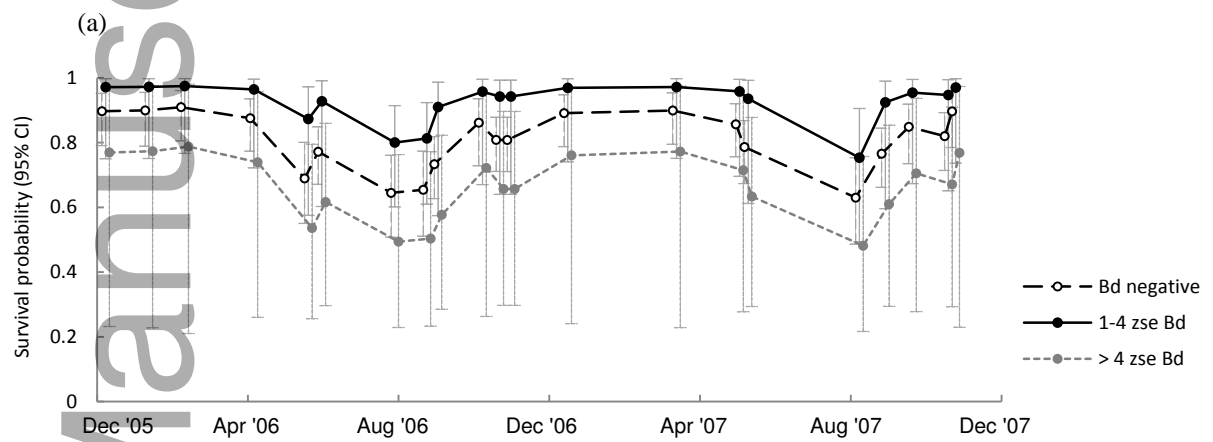


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859 **Figure 4.**

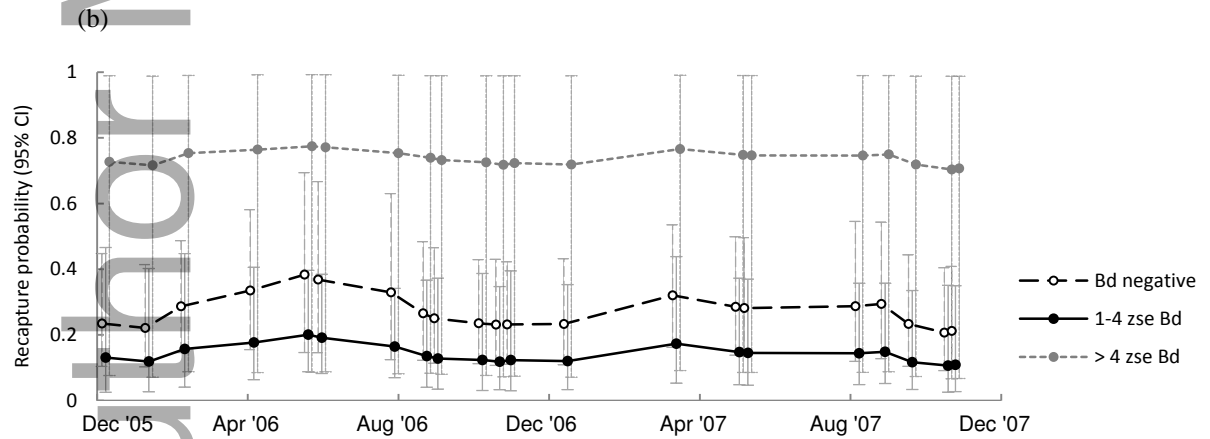
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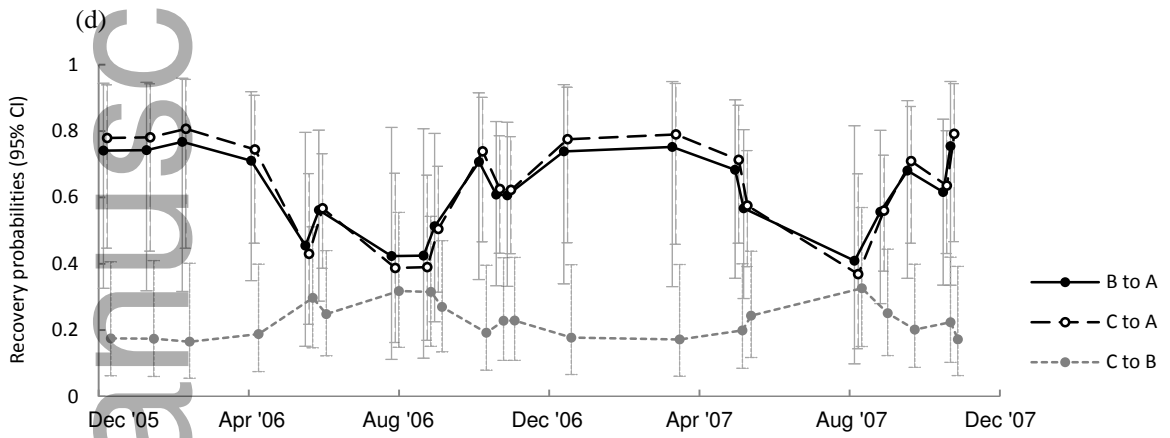
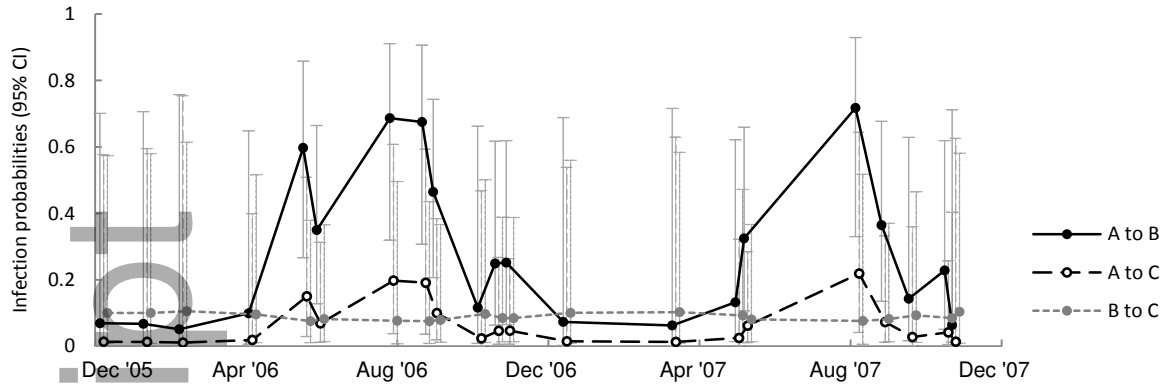


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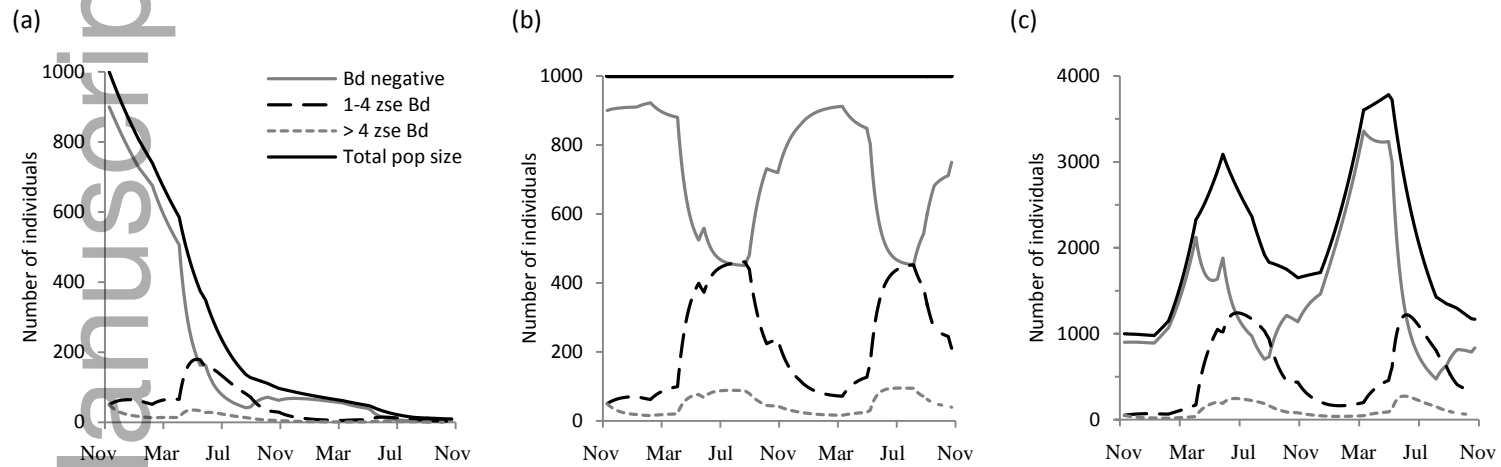
866

(c)



871 **Figure 5.**

872



873 **Supporting Information**

874

875 The following Supporting Information is available for this article online:

876 **Appendix S1.** Portable document file (PDF) containing description of parameters,
877 predictor variables, construction of candidate model sets, and population dynamics
878 simulation methods.

879 **Appendix S2.** Portable document file (PDF) containing tables of results (Tables S1-
880 S6), population dynamics simulation results, transition probabilities description, and
881 results figure (Fig. S1).

882 **Appendix S3.** Excel spreadsheet containing raw encounter history and predictor
883 variable data, together with three-state model-averaged parameter estimates.



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